

Lecture 4: Discrete Random Variables and Distributions

MSU-STT-351-Sum-19B

Contents

In this lecture, we discuss the following:

- 1 Random Variables
- 2 Discrete Random Variables
- 3 Discrete Probability Distributions
- 4 Expected Value and Variance
- 5 Homework

Definition 1

Random Variable: A random variable X is a real-valued function on the sample space S . That is, $X : S \rightarrow R$, where R is the set of all real numbers.

Note that the value of a random variable depends on a random event.

Types of a Random Variable:

- (i) A rv X is discrete if we can list its all possible values; that is, it assumes only distinct (finite or countable) values.
- (ii) A rv X is called continuous if it assumes any value in a finite or infinite interval.

Discrete Random Variables

Example 1 Suppose an **unbiased** (that is, $P(H) = P(T) = 1/2$) coin is tossed (independently) twice and the number of heads is observed. Then the sample space is

$$S = \{HH, HT, TH, TT\}.$$

Let the rv X denote the number of heads observed. Then $X : S \rightarrow R$ and $X \in \{0, 1, 2\}$. Note X is discrete. Since the coin is unbiased, its distribution can be obtained, using independence, as

$$P(X = 0) = P(TT) = P(T)P(T) = \frac{1}{4};$$

$$P(X = 1) = P(HT) + P(TH) = \frac{1}{2};$$

$$P(X = 2) = P(HH) = \frac{1}{4}.$$

Discrete Random Variables

Example 2 Suppose a coin is tossed (independently) till the first head (H) occurs. Then

$$S = \{H, TH, TTH, \dots\},$$

which is a countable set. Let the rv X denote the number of trials for the first H. Then $X \in \{1, 2, 3, \dots\} = \{n \mid n \geq 1 \text{ and an integer}\}$. Note X is again a discrete rv.

Suppose $P\{H\} = p$, $0 < p < 1$, and $P\{T\} = q$ so that $q = (1 - p)$. Then

$$P(X = 1) = P\{H\} = p,$$

$$P(X = 2) = P\{TH\} = P\{T\}.P\{H\} = qp, \text{ and}$$

$$P(X = n) = q^{n-1}p, \quad n = 1, 2, \dots$$

Example 3 (Ex.5) If the sample space S is an infinite set, does it imply any r.v. on S will also assume infinitely many values? If no, give a counter example.

Solution: No. Consider the experiment in which a coin is tossed repeatedly until a H results.

Let $Y = 1$ if the experiment terminates with at most 4 tosses and $Y = 0$ otherwise. Though the sample space is infinite, the r.v. $Y \in \{0, 1\}$.

Probability Distributions

Definition: Let X be a discrete rv with $X \in \{x_n | n \geq 1\}$. Then

(i) The probability distribution or probability mass function (*pmf*) of X is $p(x_i) = P(X = x_i)$, $x_i \in S$. Note $0 \leq p(x_i) \leq 1$, and $\sum_{x_i} p(x_i) = 1$.

It is the collection of all possible values of X and the corresponding probabilities. That is,

Values : x_1, x_2, \dots
Probabilities: $p(x_1), p(x_2), \dots$

(ii) The **cumulative distribution function** (*cdf*) of X corresponding to *pmf* $p(x)$ is

$$F(x) = P(X \leq x) = \sum_{x_i \leq x} p(x_i), \text{ for } x \in \mathbb{R}.$$

Discrete Distributions

(iii) One use of the *cdf* is in calculating probabilities. For reals a and b with $a \leq b$,

$$P(a \leq X \leq b) = F(b) - F(a-);$$

$$P(a < X \leq b) = F(b) - F(a)$$

The notation $F(a-)$ indicates evaluation of F at the point immediately to the left of a (that is, the left limit of F at a). For most discrete random variables, F changes only at integer values n so that $F(n-) = F(n - 1)$. In this case,

$$p(n) = P(X = n) = F(n) - F(n - 1), \quad n \in \mathbb{Z}_+ = \{1, 2, \dots\}.$$

Example 4: Consider the Example 2, where X denotes the number of trials required for the first head H . Then its *pmf* is

$$P(X = n) = \begin{cases} q^{n-1}p & n = 1, 2, \dots \\ 0, & \text{otherwise,} \end{cases} \quad (1)$$

and is called the geometric distribution, denoted by $G(p)$.

Here, $p = P(H)$, the probability of getting a head and $q = (1 - p)$.

As p changes, the probability value or the *pmf* changes, and p is called the **parameter** of the distribution.

Example 5 (Ex.11): An automobile service facility specializing in engine tune-ups knows that 45% of all tune-ups are done on four-cylinder automobiles, 40% on six-cylinder automobiles, and 15% on eight-cylinder automobiles. Let X = the number of cylinders on the next car to be tuned.

- What is the *pmf* of X ?
- Draw both a line graph and probability histogram for the *pmf* of part (a).
- What is the probability that the next car to be tuned has at least six cylinders? More than six cylinders?

Discrete Distributions

Solution: (a)

x:	4	6	8
p(x):	0.45	0.40	0.15

(b)

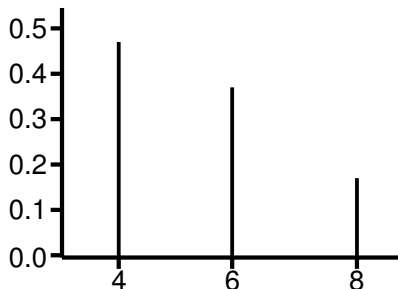


Figure: Probability Mass Function

Discrete Distributions

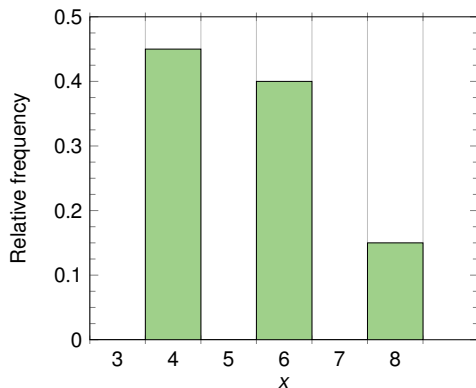


Figure: Probability Histogram

$$(c) P(X \geq 6) = P(X = 6) + P(X = 8) = 0.40 + 0.15 = 0.55.$$

Example 6 (Ex 15): Suppose a computer manufacturer receives computer boards **in lots of five**. Two boards are selected from each lot for inspection. Represent possible outcomes of the selection process by pairs. For example, the pair $(1, 2)$ represents the selection of boards 1 and 2 for inspection.

(a) List the ten different possible outcomes.

(b) Suppose that boards 1 and 2 are the only defective boards in a lot of five. Two boards are to be chosen at random. Define X to be the number of defective boards observed among those inspected. Find the probability distribution of X .

(c) Find the *cdf* $F(x)$ of X . (Hint: First determine $F(0) = P(X \leq 0)$, $F(1)$ and $F(2)$; then obtain $F(x)$ for all other x .)

Solution:

(a) Note a lot $L = \{1, 2, 3, 4, 5\}$ contain five boards. Two boards are drawn at random. The sample space is

$$S = \left\{ \{1, 2\}, \{1, 3\}, \{1, 4\}, \{1, 5\}, \{2, 3\}, \{2, 4\}, \{2, 5\}, \{3, 4\}, \{3, 5\}, \{4, 5\} \right\}.$$

(b) Note $X \in \{0, 1, 2\}$, as it denotes the number of defective boards. Then

$$P(X = 0) = p(0) = P[\{(3, 4)(3, 5)(4, 5)\}] = \frac{3}{10} = 0.3;$$

$$P(X = 2) = p(2) = P[\{(1, 2)\}] = \frac{1}{10} = 0.1;$$

$$P(X = 1) = p(1) = 1 - [p(0) + p(2)] = .60$$

and $P(X) = 0$ if $x \neq 0, 1, 2$.

Discrete Distributions

(c) Note $F(x) = P(X \leq x)$. So,

$$F(0) = P(X \leq 0) = P(X = 0) = .30$$

$$F(1) = P(X \leq 1) = P(X = 0) + P(X = 1) = .90$$

$$F(2) = P(X \leq 2) = 1.$$

Hence, the *cdf* of X is

$$F(x) = \begin{cases} 0, & x < 0 \\ 0.30, & 0 \leq x < 1 \\ 0.90, & 1 \leq x < 2 \\ 1, & 2 \leq x. \end{cases}$$

Discrete Distributions

Example 7 (Ex. 24) An insurance company offers its policyholders a number of different premium payment options. For a randomly selected policyholder, let X denote the number of months between successive payments. The *cdf* of X is as follows:

$$F(x) = \begin{cases} 0, & x < 1 \\ 0.30, & 1 \leq x < 3 \\ 0.40, & 3 \leq x < 4 \\ 0.45, & 4 \leq x < 6 \\ 0.60, & 6 \leq x < 12 \\ 1, & 12 \leq x. \end{cases}$$

- (a) What is the *pmf* of X ?
- (b) Using just the *cdf*, compute $P(3 \leq X \leq 6)$ and $P(4 \leq X)$.

Solution:

(a) Possible values of X are those values x at which $F(x)$ jumps, and $P(X = x)$ is the size of the jump at x . Thus the *pmf* of X is

$x :$	1	3	4	6	12
$p(x) :$	0.30	0.10	0.05	0.15	0.40

(b) Using *cdf*,

$$P(3 \leq X \leq 6) = F(6) - F(3^-) = F(6) - F(1) = 0.60 - 0.30 = 0.30;$$

$$P(4 \leq X) = 1 - P(X < 4) = 1 - F(4^-) = 1 - \{F(1) + F(3)\} = 1 - 0.40 = 0.60.$$

Example 8 (Ex.18) Two fair six-sided dice are tossed independently. Let M = the maximum of the two tosses (so $M(1, 5) = 5$, $M(3, 3) = 3$, etc.).

- (a) What is the *pmf* of M ?
(b) Determine the *cdf* of M and graph it.

Solution: (a): Note $M \in \{1, 2, \dots, 6\} = S$. Hence,

$$p(1) = P(M = 1) = P((1, 1)) = \frac{1}{36};$$

$$p(2) = P(M = 2) = P((1, 2) \text{ or } (2, 1) \text{ or } (2, 2)) = \frac{3}{36};$$

$$p(3) = P(M = 3) = P(1, 3) \text{ or } (2, 3) \text{ or } (3, 1) \text{ or } (3, 2) \text{ or } (3, 3)) = \frac{5}{36}.$$

Similarly, $p(4) = \frac{7}{36}$, $p(5) = \frac{9}{36}$ and $p(6) = \frac{11}{36}$.

(b) The *cdf* of X is

$$F(m) = \begin{cases} 0, & m < 1 \\ \frac{1}{36}, & 1 \leq m < 2 \\ \frac{4}{36}, & 2 \leq m < 3 \\ \frac{9}{36}, & 3 \leq m < 4 \\ \frac{16}{36}, & 4 \leq m < 5 \\ \frac{25}{36}, & 5 \leq m < 6 \\ 1, & 6 \leq m. \end{cases}$$

Discrete Distributions

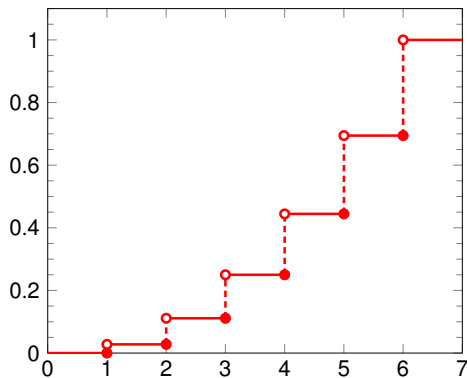


Figure: The cdf of M

Expected Value and Variance

Definition 2

Let X be a discrete random variable with *pmf* $p(x)$. The expected value or the mean value of X , denoted by $\mu_x = E(X)$, is

$$\begin{aligned}\mu_x &= x_1P(X = x_1) + x_2P(X = x_2) + \dots \\ &= \sum_{i=1}^{\infty} x_i p(x_i).\end{aligned}$$

The variance of X , denoted by $\text{Var}(X) = \sigma_x^2$ is

$$\begin{aligned}\sigma_x^2 &= E(X - \mu_x)^2 \\ &= (x_1 - \mu_x)^2 P(X = x_1) + (x_2 - \mu_x)^2 P(X = x_2) + \dots \\ &= \sum_{i=1}^{\infty} (x_i - \mu_x)^2 p(x_i).\end{aligned}$$

Expected Value and Variance

Sometimes, it is simply written as

$$\mu_x = E(X) = \sum_i x_i P(X = x_i);$$
$$\sigma_x^2 = \text{Var}(X) = \sum_i (x_i - \mu_x)^2 P(X = x_i).$$

The standard deviation (SD) of X is $\sigma_x = \sqrt{\sigma_x^2}$. (the positive square root).

Indeed, the expected value of $h(X)$, a function of X , is defined as

$$E(h(X)) = \sum_i h(x_i) p(x_i).$$

Expected Value and Variance

A shortcut formula for σ_X^2

The variance of X can be computed easily using the formula

$$\sigma_X^2 = E(X^2) - [E(X)]^2 = \sum_i x_i^2 p(x_i) - \mu_X^2.$$

Rules for Mean and the Variance

The following relations can be easily established:

- a) $\mu_{aX+b} = E(aX + b) = a\mu_X + b;$
- b) $\mu_{X+Y} = E(X + Y) = \mu_X + \mu_Y;$
- c) $\sigma_{aX+b}^2 = \text{Var}(aX + b) = a^2\sigma_X^2;$
- d) $\sigma_{aX+b} = |a|\sigma_X$

for **any** random variables X and Y .

Expected Value and Variance

Definition 3 (Independence of RVs)

The random variables X and Y are independent if

$$P(X \in A; Y \in B) = P(X \in A)P(Y \in B),$$

for all events A and B .

If X and Y are independent rvs, then

$$\sigma_{X+Y}^2 = \sigma_X^2 + \sigma_Y^2; \text{ but}$$

$$\sigma_{X+Y} = \sqrt{(\sigma_X^2 + \sigma_Y^2)} \neq \sigma_X + \sigma_Y$$

Expected Value and Variance

Example 9 (Ex 29) The *pmf* for X = the number of major defects on a randomly selected appliance of a certain type is

$x:$	0	1	2	3	4
$p(x):$	0.08	0.15	0.45	0.27	0.05

- Compute the following:
- (a) $E(X)$
 - (b) $V(X)$ directly from the definition
 - (c) The standard deviation of X
 - (d) $V(X)$ using the shortcut formula

Expected Value and Variance

Solution:

(a):

$$\begin{aligned} E(X) &= \sum_{i=0}^4 x_i p(x_i) \\ &= (0)(.08) + (1)(.15) + (2)(.45) + (3)(.27) + (4)(.05) = 2.06. \end{aligned}$$

(b):

$$\begin{aligned} V(X) &= \sum_{i=0}^4 (x_i - 2.06)^2 p(x_i) \\ &= (0 - 2.06)^2 (.08) + \dots + (4 - 2.06)^2 (.05) = .9364. \end{aligned}$$

(c): $\sigma_x = \sqrt{(0.9364)} = 0.9677.$

(d): Using the shortcut formula,

$$V(X) = \sum_{i=0}^4 x_i^2 p(x_i) - (2.06)^2 = 5.1800 - 4.2436 = 0.9364.$$

Expected Value and Variance

Example 10 (Ex 37) Suppose n candidates for a job have been ranked $1, 2, 3, \dots, n$. Let X = the rank of a randomly selected candidates, so that X has *pmf*

$$P(X = x) = p(x) = \begin{cases} \frac{1}{n}, & x = 1, 2, \dots, n \\ 0, & \text{otherwise.} \end{cases}$$

This is called the **discrete uniform distribution**.

Compute $E(X)$ and $V(X)$, using the shortcut formula. [Hint: The sum of the first n positive integers is $n(n + 1)/2$, whereas the sum of their squares is $n(n + 1)(2n + 1)/6$.]

Expected Value and Variance

Solution:

$$E(X) = \sum_{k=1}^n k \left(\frac{1}{n}\right) = \left(\frac{1}{n}\right) \sum_{k=1}^n k = \frac{1}{n} \left[\frac{n(n+1)}{2} \right] = \frac{n+1}{2},$$

$$E(X^2) = \sum_{k=1}^n k^2 \left(\frac{1}{n}\right) = \frac{1}{n} \left[\frac{n(n+1)(2n+1)}{6} \right] = \frac{(n+1)(2n+1)}{6}.$$

So,

$$\text{Var}(X) = \frac{(n+1)(2n+1)}{6} - \left(\frac{n+1}{2}\right)^2 = \frac{n^2 - 1}{12}.$$

Expected Value and Variance

Example 11: Find the expected value and SD of random variable X with

x:	0	1	2
P(X=x):	0.2	0.4	0.4

Solution:

(The mean and the variance are given by:

$$E(X) = \mu_x = 0(0.2) + 1(0.4) + 2(0.4) = 1.2$$

$$\begin{aligned} V(X) &= (0 - 1.2)^2(0.2) + (1 - 1.2)^2(0.4) + (2 - 1.2)^2(0.4) \\ &= 0.56 \end{aligned}$$

Hence,

$$\sigma_x = \sqrt{0.56} = 0.748.$$

Expected Value and Variance

Exercise 38: Let X denote the outcome when a fair die rolled once. If before the die is rolled you are offered either $(1/3.5)$ dollars or $h(X) = 1/X$ dollars, after the die is rolled, would you accept the guaranteed amount or would you gamble? [Note: It is **not** generally true that $1/E(X) = E(1/X)$]

Solution: Note that if you gamble, your average gain is

$$E[h(X)] = E\left(\frac{1}{X}\right) = \sum_{i=1}^6 \left(\frac{1}{x_i}\right) p(x_i) = \frac{1}{6} \sum_{i=1}^6 \frac{1}{x_i} = 0.408.$$

If you do not gamble, you get $\frac{1}{3.5} = 0.286$. Hence, you expect to win more if you gamble.

Expected Value and Variance

Example 12: Kids: A couple plans to have children until they get a girl, but they agree they will not have more than three children even if all are boys.

- a) Create a probability model for the number of children they'll have.
- b) Find the expected number of children.
- c) Find the expected number of boys they'll have.
- d) Find the SD of the number of children the couple may have.

Expected Value and Variance

Solution:

Outcomes	Prob.
G	0.5
BG	$(0.5)(0.5)=0.25$
BBG	$(0.5)(0.5)(0.5)=0.125$
BBB	$(0.5)(0.5)(0.5)=0.125$

Let X =Number of children. Then the *pmf* of X is

(a)

$x:$	1	2	3
$p(x):$	0.5	0.25	0.25

Expected Value and Variance

(b)

$$E(X) = 1(.5) + 2(.25) + 3(.25) = 1.75$$

(c) Let Y =Number of boys. Then the *pmf* of Y is

$y:$	0	1	2	3
$p(y):$	0.5	0.25	0.125	0.125

$$\begin{aligned}(d) : \sigma_X^2 &= (1 - 1.75)^2(.5) + (2 - 1.75)^2(.25) + (3 - 1.75)^2(.25) \\ &= 0.6875.\end{aligned}$$

Hence, $\sigma_X = \sqrt{0.6875} = 0.829$.

Expected Value and Variance

Example 13: Given **independent** random variables with means and standard deviations as shown, find the mean and standard deviation of each of these variables:

- (a) $X - 20$
- (b) $0.5Y$
- (c) $X + Y$
- (d) $X - Y$

	Mean	SD
X:	80	12
Y:	12	3

Expected Value and Variance

Solution:

$$(a) \quad \begin{aligned} \mu_{X-20} &= \mu_X - 20 = 80 - 20 = 60 \\ \sigma_{X-20} &= \sigma_X = 12. \end{aligned}$$

$$(b) \quad \begin{aligned} \mu_{(0.5)Y} &= (0.5)\mu_Y = (0.5)(12) = 6 \\ \sigma_{(0.5)Y} &= |0.5|\sigma_Y = (0.5)(3) = 1.5 \end{aligned}$$

$$(c) \quad \begin{aligned} \mu_{X+Y} &= \mu_X + \mu_Y = 80 + 12 = 92 \\ \sigma_{X+Y} &= \sqrt{\sigma_X^2 + \sigma_Y^2} = \sqrt{12^2 + 3^2} = \sqrt{153} = 12.369 \end{aligned}$$

$$(d) \quad \begin{aligned} \mu_{X-Y} &= \mu_X - \mu_Y = 80 - 12 = 68 \\ \sigma_{X-Y} &= \sqrt{\sigma_X^2 + \sigma_{-Y}^2} \\ &= \sqrt{\sigma_X^2 + (-1)^2\sigma_Y^2} \\ &= \sqrt{\sigma_X^2 + \sigma_Y^2} \\ &= 12.369. \end{aligned}$$

Homework

Homework:

Section 3.1: 1, 7, 9

Section 3.2: 11, 13, 18, 23

Section 3.3: 33, 35, 39, 43

-- END OF LECTURE 4 --