Preliminary Exam: Probability

9:00 am to 2:00 pm, August 25, 2006

Problem 1. Assume that triangular array  $\{X_{n,k}: k=1,...,n\}$  is independent for each n=1,...

(a) Prove:  $\max_{1 \le k \le n} \{ |X_{n,k}| \} \xrightarrow[n \to \infty]{} 0$  in probability **if and only if** 

$$\sum_{k=1}^{n} P(|X_{n,k}| \ge \varepsilon) \xrightarrow[n \to \infty]{} 0, \ \forall \varepsilon > 0.$$

Hint:  $1 > e^{-x} > 1 - x$ , x > 0.

(b) Let  $\{X_k, k=1,...\}$  be independent and identically distributed with

$$\frac{P(|X| > x)}{x^{-\alpha}} \xrightarrow[x \to \infty]{} C \text{ with } 0 < C < \infty \text{ for some } \alpha > 0. \text{ Let } \{a_n\} \text{ be a}$$

positive sequence. Prove:  $\frac{\max\{|X_k|\}}{n^{1/\alpha}a_n} \xrightarrow[n\to\infty]{} 0 \text{ in probability if and only}$  if  $a_n\to\infty$ .

(c) Let  $\{X_k, k=1,...\}$  be independent and identically distributed with

$$X_k \sim N(0,1)$$
. Find the minimal  $\beta > 0$  so that  $\frac{\max_{1 \le k \le n} \{ |\frac{1}{X_k}| \}}{n^{\beta + \varepsilon}} \xrightarrow[n \to \infty]{} 0$  in probability, for each  $\varepsilon > 0$ .

Problem 2. Let X be a random variable with  $P(X \ge x) = x^{-\alpha}$ ,  $x \ge 1$  for some  $\alpha > 0$  (so  $X \ge 1$  a.s.).

- (a) Calculate the function  $f(y, \alpha) = E(X^2; X \le y), \alpha > 0, y > 0$ .
- (b) For each  $\alpha > 0$  find  $\lim_{y \to \infty} \frac{y^2 P(X \ge y)}{f(y, \alpha)}$
- (c) Find an example of a sequence of **positive** random variables,  $\{X_k, k=1,...\}$ , so that the following 2 requirements both hold.

- (i) For each  $\alpha > 0$  there exist an integer  $K_{\alpha} > 0$  so that  $E[(X_k)^{\alpha}] = \infty$ ,  $k > K_{\alpha}$
- (ii)  $\{X_k\}$  converges in probability to 0.

Problem 3. In what follows a characteristic function (c.f.) of a random variable X is denoted by  $\varphi_X(t)$ .

(a) Prove that for every T > 0:

(i) 
$$\frac{1}{T} \int_{-T}^{T} \varphi_X(t) dt = 2 \cdot E(\frac{\sin(TX)}{TX})$$

(ii) 
$$E(1 - \frac{\sin(TX)}{TX}) \ge \frac{1}{2}P(|X| \ge \frac{2}{T})$$

Let  $\{X_k, k=1,...\}$  be a sequence of random variables. For parts (b) and (c) assume that  $\lim_{k\to\infty} \varphi_{X_k}(t)$  exists for |t|<1. Denote the limit by g(t) (observe that g(t) is defined only on |t|<1.)

(b) Prove that for every M > 2 we have

$$M \cdot \int_{-2/M}^{2/M} [1 - g(t)] dt \ge \limsup_{k \to \infty} \{ P(|X_k| \ge M) \}$$

- (c) Assume also that:  $\lim_{t\to 0} g(t) = 1$ .
- (i) Use part (b) to show that  $\forall \varepsilon > 0 \ \exists M > 0$  so that  $\sup_{k} \{P(|X_k| \ge M)\} \le \varepsilon$ . (i.e. :  $\{X_k\}$  is *tight*).
- (ii) Use (c)(i) to show that there exists a random variable Y whose c.f.  $\varphi_Y(t)$  is an extension of g(t) (i.e.  $\varphi_Y(t) = g(t)$ , |t| < 1.)

Problem 4. Let  $\{D_n, F_n; n=1,...\}$  be a sequence of  $L^2$  martingale differences, namely  $\{D_n\}$  are random variables and  $\{F_n\}$  are  $\sigma$  – algebras,  $D_n \in F_n, \ F_n \subset F_{n+1}, \ \mathrm{E}_{F_n}(D_{n+1}) = 0$  and  $\mathrm{E}(D_n^2) < \infty$ . We denote  $X_n \equiv \sum_{k=1}^n D_k$ ,  $A_n \equiv \sum_{k=1}^n \mathrm{E}_{F_{k-1}}(D_k^2)$  ( $F_0$  is a trivial  $\sigma$  – algebra) and  $A_\infty \equiv \lim_n A_n$  (can be  $\infty$ ).

- (a) Is  $N_C = \inf\{n: A_{n+1} > C\}$ , C > 0 a stopping time? Prove or give a counter example.
- (b) Does  $\lim_{n} X_{n \wedge N_C}$  exists a.s. for each C > 0? Prove or disprove.
- (c) (i) Does  $\lim_{n} X_n$  exists a.s. on  $\bigcup_{C>0} \{N_C = \infty\}$ ? Explain.
- (ii) What is the relationship between the events  $\bigcup_{C>0} \{N_C = \infty\}$  and  $\{A_\infty = \infty\}$ ?

Problem 5. Here we use the setup and notations of problem 4.

- (a) Let  $b_n \uparrow \infty$ . Prove that on the event  $\left\{\sum_{k=1}^{\infty} \mathbb{E}_{F_{k-1}} \left(\frac{D_k^2}{b_k^2}\right) < \infty\right\}$  we have  $\frac{X_n}{b_n} \xrightarrow[n \to \infty]{} 0$ , a.s.
- (b) Prove that  $\sum_{k=2}^{\infty} E_{F_{k-1}} (\frac{D_k^2}{A_k^2}) \le \int_{A_1}^{\infty} \frac{dt}{t^2}$ , a.s.
- (c) Prove (i) Assume  $E(D_1^2) > 0$ . Prove that  $\sum_{k=1}^{\infty} \frac{D_k}{A_k}$  converges a.s.
  - (ii)  $\frac{X_n}{A_n} \xrightarrow[n \to \infty]{} 0$ , a.s. on the event  $\{A_\infty = \infty\}$
  - (iii)  $\frac{X_n}{A_n}$  converges a.s.

Problem 6.  $\{X_n\}$  is called Uniformly Integrable (UI) if  $\sup E(|X_n|\cdot 1_{\{|X_n|>M\}}) \xrightarrow[M\to\infty]{} 0$ .

- (a) Prove:  $\{X_n\}$  is UI if and only if
- (i)  $\sup_{n} E(|X_n|) < \infty$ , and

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- (ii)  $\forall \varepsilon > 0 \ \exists \delta > 0 \ \text{so that} \ P(A) < \delta \Rightarrow \sup_{n} E(|X_n|; A) \le \varepsilon$
- (b) Give an example of a sequence  $\{X_n\}$  that is not UI but at the same time the following 3 requirements hold: (i)  $X_n \xrightarrow[n \to \infty]{} 0$ , a.s.,

(ii) 
$$E(X_n) \xrightarrow[n \to \infty]{} 0$$
, (iii)  $\sup_n E(|X_n|) < \infty$ .

- (c) Give an example of a sequence  $\{X_n\}$  so that (i)  $\{X_n\}$  is UI
- (ii)  $X_n \xrightarrow[n \to \infty]{} 0$ , a.s. (iii) There exist a  $\sigma$ -algebra F so that  $E_F(X_n)$  doesn't converge a.s.

Problem 7. Let  $\{X_k, k=1,...\}$  be independent and identically distributed with **symmetric** distribution. Let  $a_0=0$ ,  $a_n \uparrow \infty$  and denote  $Y_n \equiv X_n \cdot 1_{\{|X_n| \le a_n\}}$ . Assume

- (i) There is C > 0 so that  $\sum_{n=m}^{\infty} a_n^{-2} \le C \cdot m \cdot a_m^{-2}$ ,  $\forall m \ge 1$ , and
- (ii)  $\sum_{n=1}^{\infty} P(|X_1| \ge a_n) < \infty$
- (a) Prove:  $\sum_{n=1}^{\infty} \frac{E(Y_n^2)}{a_n^2} < \infty$ . Hint:  $E(Y_n^2) = \sum_{m=1}^n E(X_m^2; a_{m-1} < |X_m| \le a_m)$ , etc.
- (b) Prove:  $\frac{\sum_{k=1}^{n} X_{k}}{a_{n}} \xrightarrow[n \to \infty]{} 0, \text{ a.s.}$

(c) Assume  $E|X_1|^p < \infty$ , 1 . Show by using part (b) that

$$\frac{\sum_{k=1}^{n} X_k}{n^{1/p}} \xrightarrow[n \to \infty]{} 0, \text{ a.s.}$$

Problem 8. Let  $\{B(t)\}$  denote standard Brownian motion.

(a) (i) Prove that 
$$\frac{1}{\sqrt{2\pi}} \int_{x}^{\infty} e^{-z^2/2} dz \le \frac{1}{2} \cdot e^{-x^2/2}, \ x > 0$$
 (Hint: use  $z = x + y$ ).

Remark: in the book the upper bound is  $\frac{e^{-x^2/2}}{x}$  which is worse for x < 2.

- (ii) Use the inequality in (i) to present a function f(t, x) so that  $P(\max_{0 < u < t} | B_u | > x) \le f(t, x), x > 0.$
- (b) Prove that  $E \max_{0 \le u \le t} B_u^2 \le 2 \cdot t$ .
- (c) Let  $\Delta_n = \max\{\Delta_{m,n}: 1 \le m \le 2^n\}$  where

$$\Delta_{m,n} = \max\{|B(t) - B(\frac{m-1}{2^n})\}: \frac{m-1}{2^n} \le t < \frac{m}{2^n}\}.$$
 Prove: There is  $C < \infty$  so that

$$\Delta_n \leq C \cdot \sqrt{n \cdot 2^{-n}}, \ n \geq N(\omega)$$